Lecture 13

1 Variance of discrete random variables

We have already encountered *standard deviation* when we talked about the normal density. It was a measure of how much the distribution deviates from the mean. It is now time to get a more formal grasp on that quantity.

The square of standard deviation is called *variance*, and just like standard deviation is usually written as σ , variance is often written as σ^2 . It is often much more convenient to talk about variance than about standard deviation. The chief reason for that is, that variance of a sum of two **independent** random variables is the sum of their variances. This is clearly not true of standard deviation, since it must remain a square root of the variance.

Definition 1 Let X be a numerically valued random variable with expected value $\mu = E(X)$. Then the variance of X, here denoted by V(X), is

$$V(X) = E((X - \mu)^2).$$

Example 1 Let X be the result of rolling a six-sided die. Then E(X) = 3.5 and:

$$V(X) = E((X-3.5)^2) = \sum_{x=1}^{6} (x-3.5)^2 \times \frac{1}{6} = \frac{1}{6} ((1-3.5)^2 + (2-3.5)^2 + (3-3.5)^2 + (4-3.5)^2 + (5-3.5)^2 + (6-3.5)^2)$$

 $=\frac{1}{6}(\frac{25}{4}+\frac{9}{4}+\frac{1}{4}+\frac{1}{4}+\frac{9}{4}+\frac{25}{4})=\frac{70}{24}=\frac{35}{12}$

Then the standard deviation is $\sigma = \sqrt{35/12}$.

Theorem 1 For any random variable X with $E(X) = \mu$,

$$V(X) = E(X^2) - \mu^2.$$

Proof By definition, $V(X) = E((X - \mu)^2)$. Then:

$$V(X) = E((X - \mu)^2) = E(X^2 - 2\mu X + \mu^2) = E(X^2) - 2\mu E(X) + \mu^2,$$

since μ is a numerical quantity. Further, $E(X) = \mu$ and so:

$$V(X) = E(X^{2}) - 2\mu^{2} + \mu^{2} = E(X) - \mu^{2}.$$

Therefore, there is an easier way to compute the variance of a roll of a die:

$$E(X^2) = \frac{1}{6}(1+4+9+16+25+36) = \frac{91}{6}$$

$$\mu^2 = \frac{49}{4}$$

And so:

$$V(X) = \frac{182 - 147}{12} = \frac{35}{12}.$$

2 Worksheet problems

- Knowing that V(X) = E(X²) μ², find simple expressions for V(cX) and V(X + c), for a constant c.
 If X and Y are independent random variables, then V(X + Y) = V(X) + V(Y). What is V(X) if X is the number of successes out of n Bernoulli trials with probability p?

$$V(cX) = c^2 V(X)$$
$$V(X+c) = V(X)$$

$$\begin{split} V(cX) &= E((cX)^2) - (c\mu)^2 = c^2(E(X^2 - \mu^2)) = c^2V(X) \\ V(X+c) &= E((X+c)^2) - (\mu+c)^2 = E(X^2 + 2cX + c^2) - (\mu^2 + 2\mu c + c^2) \\ &= E(X^2) + 2cE(X) + c^2 - \mu^2 - 2\mu c - c^2 = E(X^2) - \mu^2 = V(X) \end{split}$$

If X and Y are independent random variables,

$$V(X+Y) = V(X) + V(Y).$$

Example 2 (Bernoulli trials) Bernoulli trials are independent events. The variance of a single trial is:

$$E(X^{2}) - \mu^{2} = p - p^{2} = p(1 - p)$$

There are n of them, so the combined variance is:

np(1-p) = npq

3 Expectation and variance of continuous random variables

For continuous random variables, expectation and variance work in a way analogous to the discrete case.

Definition 2 Let X be a real-valued random variable with density function f(x). The expected value $\mu = E(X)$ is defined by:

$$\mu = E(X) = \int_{-\infty}^{\infty} x f(x) dx,$$

provided that the integral $\int_{-\infty}^{\infty} |x| f(x) dx$ is finite.

For any real-valued random variables X, Y with E(X), E(Y):

$$E(X + Y) = E(X) + E(Y)$$
$$E(cX) = cE(X)$$

Example 3 (Uniform Density) recall from last week's homework, that for uniform density on (a, b), f(x) = 1/(b-a):

$$\mu = \frac{1}{b-a} \int_{a}^{b} x dx = \frac{1}{b-a} \left[\frac{x^2}{2} \right]_{a}^{b} = \frac{b^2 - a^2}{2(b-a)} = \frac{b+a}{2}$$

Example 4 (Exponential Density) Similarly, for exponential density $f(x) = \lambda e^{-\lambda x}$ on $(0, \infty)$:

$$\mu = \int_0^\infty x f(x) dx = \int_0^\infty x \lambda e^{-\lambda x} dx = \left[\frac{-e^{-\lambda x}}{\lambda} - xe^{-\lambda x}\right]_0^\infty = \frac{1}{\lambda}$$

If X, Y are mutually independent:

$$E(X \times Y) = E(X) \times E(Y)$$

Definition 3 Let X be a real-valued random variable with density function f(x) and expected value $\mu = E(X)$. The variance $\sigma^2 = V(X)$ is:

$$\sigma^2 = V(X) = E((X - \mu)^2) = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

Just as in the discrete case:

$$V(X) = E(X^2) - \mu^2$$
$$V(cX) = c^2 V(X)$$
$$V(X + c) = V(X)$$

And if X, Y are independent:

$$V(X+Y) = V(X) + V(Y)$$

4 Expected number of trials until first success

Suppose that we are going to flip a coin until the first time it comes up "tails." Each time we have 1/2 chance of success. Let N be a random variable equal to the total number of flips. What is E(N)?

Here is a brute force calculation:

$$E(n) = \sum_{i=1}^{\infty} \frac{i}{2^{i}} = \frac{1}{2} + \frac{2}{4} + \frac{3}{8} + \dots = \frac{1}{2} + \frac{1}{4} + \frac{1}{4} + \frac{1}{8} + \frac{1}{8} + \frac{1}{8} + \frac{1}{8} + \dots$$
$$= \underbrace{\frac{1}{2} + \frac{1}{4} + \frac{1}{8} + \dots}_{\sum_{i=1}^{\infty} \frac{1}{2^{i}}} + \underbrace{\frac{1}{4} + \frac{1}{8} + \dots}_{\sum_{i=2}^{\infty} \frac{1}{2^{i}}} + \underbrace{\frac{1}{8} + \dots}_{\sum_{i=1}^{\infty} \frac{1}{2^{i}}} + \underbrace{\frac{1}{8} + \dots}_{\max}_{\sum_{$$

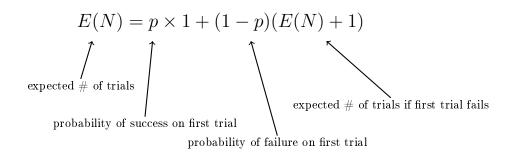
However, we can argue the following. When we flip the coin for the first time, with probability 1/2 it comes up tails. So in the sum that determines expected value, 1 has weight 1/2. Now, with probability 1/2 the flip is heads, so having made one flip we are exactly where we started, and have gained nothing - the expected number of flips FROM THIS POINT ON is exactly the same as the expected number of flips at the beginning. So:

$$E(N) = 1/2 + 1/2(1 + E(N))$$

we can multiply both sides by 2 and get:

$$2E(N) = 1 + 1 + E(N)$$
$$E(N) = 2$$

In general, if the chance of success on each trial is p, the expected number of trials until the first success is:



Now, here's a another way to get the same result. We have:

$$E(N) = \sum_{k=1}^{\infty} k(1-p)^{k-1} p = p \sum_{k=1}^{\infty} k(1-p)^{k-1} = p \sum_{k=1}^{\infty} kq^{k-1}$$

We know that the sum of a geometric sequence works as follows:

$$\sum_{k=0}^{\infty} q^k = 1 + \sum_{k=1}^{\infty} q^k = \frac{1}{1-q}$$

To get the term in the middle it's enough to pull out the first term of the sum. Now, we can differentiate both sides of this equation:

$$\left(1 + \sum_{k=1}^{\infty} q^k\right)' = \left(\frac{1}{1-q}\right)'$$
$$\sum_{k=1}^{\infty} kq^{k-1} = \frac{1}{(1-q)^2} = \frac{1}{p^2}$$

And going back to the expression for E(N), we get:

$$E(N) = p \sum_{k=1}^{\infty} k(1-p)^{k-1} = \frac{p}{p^2} = \frac{1}{p}$$

5 Coupon Collector's Problem

Alice would like to collect n different stickers from chips packets. In each packet, each of the n stickes can be found with probability 1/n. How many packets, on average, does she need to buy to collect all stickers?

If n = 2, the first packet definitely contains a sticker she doesn't already have. Then, each next packet contains the other sticker with probability 1/2. So the expected number of trials until first success is 2. We need 1 trial to get the first sticker, and another 2 to get the next. E(X) = 1 + 2 = 3

If n = 3, again, we get the first sticker on first trial. Then, with probability 2/3 we get a different sticker. So it takes, on average 1.5 trials to get the second sticker. Then it takes another 3 to get the third. E(X) = 1 + 1.5 + 3

In general, it takes 1 trial to get the first sticker, n/(n-1) to get the second, n/(n-2) to get the third, etc, until once we have n-1 distinct stickers it takes another n trials to get the last.

$$E(X) = 1 + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{2} + n$$

This is an incredibly important problem in computer science, since many algorithms collect things at random and this prblem describes their expected runtime.